

Statistics

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Summary

- ggplot() specifies what data to use and what variables will be mapped to where **·**
- \cdot inside $ggplot()$, $aes(x = , y = , color =) specify what variables$ correspond to what aspects of the plot in general
- layers of plots can be combined using the + at the **end** of lines **·**
- use geom_line() and geom_point() to add lines and points **·**
- sometimes you need to add a group element to aes() if your plot looks strange **·**
- make sure you are plotting what you think you are by checking the numbers! **·**
- facet_grid(~variable) and facet_wrap(~variable) can be helpful to quickly split up your plot **·**

Summary

- the factor class allows us to have a different order from alphanumeric for **·** categorical data
- we can change data to be a factor variable using mutate(), as_factor() (in **·** the forcats package), or factor() functions and specifying the levels with the levels argument
- fct_reorder({variable_to_reorder}, {variable_to_order_by}) helps **·** us reorder a variable by the values of another variable
- arranging, tabulating, and plotting the data will reflect the new order **·**

Overview

We will cover how to use R to compute some of basic statistics and fit some basic statistical models.

- Correlation **·**
- T-test **·**
- Linear Regression / Logistic Regression **·**

Overview

We will focus on how to use R software to do these. We will be glossing over the statistical **theory** and "formulas" for these tests. Moreover, we do not claim the data we use for demonstration meet **assumptions** of the methods.

There are plenty of resources online for learning more about these methods.

Check out www.opencasestudies.org for deeper dives on some of the concepts covered here and the [resource page](https://daseh.org/resources.html) for more resources.

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The correlation coefficient is a summary statistic that measures the strength of a linear relationship between two numeric variables.

- The strength of the relationship based on how well the points form a line **·**
- The direction of the relationship based on if the points progress upward or downward **·**

See this [case study](https://www.opencasestudies.org/ocs-bp-co2-emissions/#Data_Analysis) for more information.

Function cor() computes correlation in R.

```
cor(x, y = NULL, use = c("everything"
,
"complete.obs"),
    method = c("pearson"
,
"kendall"
,
"spearman"))
```
- provide two numeric vectors of the same length (arguments x, y), or **·**
- provide a data.frame / tibble with numeric columns only **·**
- by default, Pearson correlation coefficient is computed **·**

Correlation test

Function cor.test() also computes correlation and tests for association.

cor.test(x, y = NULL, alternative(c("two.sided" , "less" , "greater")), method = c("pearson" , "kendall" , "spearman"))

- provide two numeric vectors of the same length (arguments x, y), or **·**
- provide a data.frame / tibble with numeric columns only **·**
- by default, Pearson correlation coefficient is computed **·**
- alternative values: **·**
	- two.sided means true correlation coefficient is not equal to zero (default) **-**
	- greater means true correlation coefficient is > 0 (positive relationship) **-**
	- less means true correlation coefficient is < 0 (negative relationship) **-**

https://daseh.org/data/Yearly_CO2_Emissions_1000_tonnes.csv

library(dasehr)

head(yearly_co2_emissions)

Correlation for two vectors

First, we compute correlation by providing two vectors.

x and y must be numeric vectors y1980 <- yearly_co2_emissions %>% pull(`1980`) $y1985 < -yearly.co2$ emissions %>% pull(`1985`)

Like other functions, if there are NAs, you get NA as the result. But if you specify use only the complete observations, then it will give you correlation using the non-missing data.

cor(y1980, y1985, use = "complete.obs")

[1] 0.9936257

Correlation coefficient calculation and test

cor.test(y1980, y1985)

Pearson's product-moment correlation

```
data: y1980 and y1985
t = 114.59, df = 169, p-value < 0.00000000000000022
alternative hypothesis: true correlation is not equal to 0
95 percent confidence interval:
 0.9913844 0.9952853
sample estimates:
      cor
0.9936257
```
Broompackage

The broom package helps make stats results look tidy

library(broom) cor_result <- tidy(cor.test(y1980, y1985)) glimpse(cor_result)

Rows: 1 Columns: 8 \$ estimate <dbl> 0.9936257 \$ statistic <dbl> 114.5851 \$ p.value <dbl> 0.00 \$ parameter <int> 169 \$ p.value <dbl> 0.0000000
\$ parameter <int> 169
\$ conf.low <dbl> 0.9913844 \$ conf.high <dbl> 0.9952853 \$ method <chr> "Pearson's product-moment correlation" \$ alternative <chr> "two.sided"

Correlation for two vectors with plot

In plot form… geom_smooth() and annotate() can help.

```
corr_value < -pull(cor_result, estimate) % round(digits = 4)
cor_label <- paste0("R = ", corr_value)
yearly_co2_emissions %>%
  ggplot(aes(x = '1980', y = '1985')) + geom\_point(size = 1) + geom\_smooth() +annotate("text", x = 2000000, y = 4000000, label = cor_label)
```


Correlation for data frame columns

We can compute correlation for all pairs of columns of a data frame / matrix. This is often called, *"computing a correlation matrix"*.

Columns must be all numeric!

```
co2_subset <- yearly_co2_emissions %>%
  select(c(`1950`, `1980`, `1985`, `2010`))
head(co2_subset)
\# \wedge tibble: 6 \times 4
```


Correlation for data frame columns

We can compute correlation for all pairs of columns of a data frame / matrix. This is often called, *"computing a correlation matrix"*.

cor_mat <- cor(co2_subset, use = "complete.obs") cor mat

 1980 1985 2010 1.0000000 0.9228253 0.8818288 0.5415047 0.9228253 1.0000000 0.9935477 0.7270839 0.8818288 0.9935477 1.0000000 0.7827256 0.5415047 0.7270839 0.7827256 1.0000000

Correlation for data frame columns with plot

corrplot package can make correlation matrix plots

library(corrplot) corrplot(cor_mat)

Correlation does not imply causation

[source](http://doi.org/10.1007/s10393-020-01472-1)

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T-test

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T-test

The commonly used are:

- **one-sample t-test** used to test mean of a variable in one group **·**
- **two-sample t-test** used to test difference in means of a variable between **·** two groups (if the "two groups" are data of the *same* individuals collected at 2 time points, we say it is two-sample paired t-test)

The **t** . **test()** function in R is one to address the above.

```
t.test(x, y = NULL,
       alternative = c("two.sided"
,
"less"
,
"greater"),
       mu = 0, paired = FALSE, var.equal = FALSE,
       conf. level = 0.95, ...
```
Runningone-samplet-test

It tests the mean of a variable in one group. By default (i.e., without us explicitly specifying values of other arguments):

- \cdot tests whether a mean of a variable is equal to 0 (mu = θ)
- uses "two sided" alternative (alternative = "two.sided") **·**
- returns result assuming confidence level 0.95 (conf.level = 0.95)
- omits NA values in data **·**

Let's look at the CO2 emissions data again.

t.test(y1980)

One Sample t-test

```
data: y1980
t = 3.3324, df = 170, p-value = 0.001056
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval:
   44745.81 174792.25
sample estimates:
mean of x 
    109769
```
Running two-sample t-test

It tests the difference in means of a variable between two groups. By default:

- \cdot tests whether difference in means of a variable is equal to 0 (mu = θ)
- uses "two sided" alternative (alternative = "two.sided") **·**
- returns result assuming confidence level 0.95 (conf.level = 0.95) **·**
- assumes data are not paired ($paired = FALSE$) **·**
- assumes true variance in the two groups is not equal (var.equal = FALSE)
- omits NA values in data **·**

Check out this this case [study](https://www.opencasestudies.org/ocs-bp-rural-and-urban-obesity/#Data_Analysis) and this case [study](https://www.opencasestudies.org/ocs-bp-diet/#Data_Analysis) for more information.

Running two-sample t-test in R

t.test(y1980, y1985)

Welch Two Sample t-test

```
data: y1980 and y1985
t = -0.090533, df = 341, p-value = 0.9279
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 -95902.79 87462.97
sample estimates:
mean of x mean of y
 109769.0 113988.9
```
T-test: retrieving information from the result with broom package

The broom package has a tidy() function that can organize results into a data frame so that they are easily manipulated (or nicely printed)

result <- t.test(y1980, y1985) result_tidy <- tidy(result) glimpse(result_tidy)

Rows: 1 Columns: 10
\$ estimate $<$ dbl > -4219.909 \$ estimate1 <dbl> 109769 \$ estimate2 <dbl> 113988.9 \$ statistic <dbl> -0.09053303 p.value <dbl> 0.9279168
parameter <dbl> 340.999 \$ parameter <dbl> 340.999 \$ conf.low <dbl> -95902.79 \$ conf.high <dbl> 87462.97 \$ method <chr> "Welch Two Sample t-test" \$ alternative <chr> "two.sided"

P-value adjustment

You run an increased risk of Type I errors (a "false positive") when multiple hypotheses are tested simultaneously.

Use the p adjust() function on a vector of p values. Use method $=$ to specify the adjustment method:

```
my_pvalues <- c(0.049, 0.001, 0.31, 0.00001)
p.adjust(my_pvalues, method = "BH") # Benjamini Hochberg
```
[1] 0.06533333 0.00200000 0.31000000 0.00004000

```
p.adjust(my_pvalues, method = "bonferroni") # multiply by number of tests
```
[1] 0.19600 0.00400 1.00000 0.00004

my_pvalues * 4

[1] 0.19600 0.00400 1.24000 0.00004

See [here](https://www.nature.com/articles/nbt1209-1135) for more about multiple testing correction. Bonferroni also often done as p value threshold divided by number of tests (0.05/test number).

Some other statistical tests

- wilcox.test() Wilcoxon signed rank test, Wilcoxon rank sum test **·**
- shapiro.test() Shapiro test **·**
- ks.test() Kolmogorov-Smirnov test **·**
- var.test()– Fisher's F-Test **·**
- chisq.test() Chi-squared test **·**
- aov() Analysis of Variance (ANOVA) **·**

Summary

- Use cor() to calculate correlation between two vectors, cor.test() can give **·** more information.
- corrplot() is nice for a quick visualization! **·**
- **t.test()** one sample test to test the difference in mean of a single vector from zero (one input)
- **t.test()** two sample test to test the difference in means between two vectors (two inputs)
- tidy() in the broom package is useful for organizing and saving statistical test **·** output
- Remember to adjust p-values with p.adjust() when doing multiple tests on data **·**

Lab Part 1

[Class Website](https://daseh.org/)

[Lab](https://daseh.org/modules/Statistics/lab/Statistics_Lab.Rmd)

Regression

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Linear regression

Linear regression is a method to model the relationship between a response and one or more explanatory variables.

Most commonly used statistical tests are actually specialized regressions, including the two sample t-test, [see here for more](https://www.opencasestudies.org/ocs-bp-diet/#(t)-test_and_linear_regression).

Linear regression notation

Here is some of the notation, so it is easier to understand the commands/results.

$$
y_i = \alpha + \beta x_i + \varepsilon_i
$$

where:

- \cdot y_i is the outcome for person i
- *α* is the intercept **·**
- *β* is the slope (also called a coefficient) the mean change in y that we would **·** expect for one unit change in x ("rise over run")
- \cdot x_i is the predictor for person i
- **·** ε_i is the residual variation for person i

Linear regression

Linear regression

Linear regression is a method to model the relationship between a response and one or more explanatory variables.

We provide a little notation here so some of the commands are easier to put in the proper context.

$$
y_i = \alpha + \beta_1 x_{i1} + \beta_2 x_{i2} + \beta_3 x_{i3} + \varepsilon_i
$$

where:

- \cdot y_i is the outcome for person i
- *α* is the intercept **·**
- \cdot β_1 , β_2 , β_2 are the slopes/coefficients for variables x_{i1} , x_{i2} , x_{i3} average difference in y for a unit change (or each value) in x while accounting for other variables
- x_{i1} , x_{i2} , x_{i3} are the predictors for person i
- \cdot ε _{*i*} is the residual variation for person i

See this [case study](https://www.opencasestudies.org/ocs-bp-diet/#Data_Analysis) for more details.

Linear regression fit in R

To fit regression models in R, we use the function glm() (Generalized Linear Model).

You may also see lm() which is a more limited function that only allows for normally/Gaussian distributed error terms (aka typical linear regressions).

We typically provide two arguments:

- formula model formula written using names of columns in our data **·**
- data our data frame **·**

Linear regression fit in R: model formula

Model formula

$$
y_i = \alpha + \beta x_i + \varepsilon_i
$$

In R translates to

 $y - x$

Linear regression fit in R: model formula

Model formula

 $y_i = \alpha + \beta x_i + \varepsilon_i$

In R translates to

 $V - X$

In practice, y and x are replaced with the **names of columns from our data set**.

For example, if we want to fit a regression model where outcome is income and predictor is years_of_education, our formula would be:

income ~ years_of_education

Linear regression fit in R: model formula

Model formula

$$
y_i = \alpha + \beta_1 x_{i1} + \beta_2 x_{i2} + \beta_3 x_{i3} + \varepsilon_i
$$

In R translates to

$$
y - x1 + x2 + x3
$$

In practice, y and x1, x2, x3 are replaced with the **names of columns from our data set**.

For example, if we want to fit a regression model where outcome is income and predictors are years_of_education, age, and location then our formula would be:

```
income ~ years_of_education + age + location
```
Linear regression

We will use our the calenviroscreen dataset from the dasehr package to examine how traffic estimates predict diesel particulate emissions.

Linear regression: model fitting

For this model, we will use two variables:

- **DieselPM** estimated diesel particulate emissions from on-road and non-road sources **·**
- **TrafficPctl** percentile ranking of traffic density **·**

```
fit <- glm(DieselPM ~ TrafficPctl, data = calenviroscreen)
fit
```
Call: glm(formula = DieselPM ~ TrafficPctl, data = calenviroscreen)

Coefficients: (Intercept) TrafficPctl 0.042452 0.003637

Degrees of Freedom: 7999 Total (i.e. Null); 7998 Residual (35 observations deleted due to missingness) Null Deviance: 537.2 Residual Deviance: 449.1 AIC: -330.9

Linear regression: model summary

The summary() function returns a list that shows us some more detail

summary(fit)

Call: $glm(formula = DieselPM - TrafficPctl, data = calenviroscreen)$ Coefficients: Estimate Std. Error t value Pr(>|t|) $(Intercept) 0.04245151 0.00529915 8.011 0.000000000000013 **$ TrafficPctl 0.00363651 0.00009177 39.625 < 0.0000000000000002 *** --- Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 (Dispersion parameter for gaussian family taken to be 0.05614936) Null deviance: 537.24 on 7999 degrees of freedom Residual deviance: 449.08 on 7998 degrees of freedom (35 observations deleted due to missingness) AIC: -330.9

Number of Fisher Scoring iterations: 2

tidy results

The broom package can help us here too!

The estimate is the coefficient or slope.

for one change in the traffic percentile, we see 0.003637 more Diesel particulate emissions. The error for this estimate is relatively small at 0.00009. This relationship appears to be significant with a small p value < 2e-16.

tidy(fit) %>% glimpse()

```
Rows: 2
Columns: 5
$ term <chr> "(Intercept)"
,
"TrafficPctl"
$ estimate <dbl> 0.042451513, 0.003636512
$ std.error <dbl> 0.00529915058, 0.00009177366
$ statistic <dbl> 8.011003, 39.624789
$ p.value <dbl> 0.0000000000000012983396857238743156732797112774494962140342
```
Linear regression: multiple predictors

Let's try adding another explanatory variable to our model, amount of daily Ozone concentration (Ozone). Ozone is usually inversely related to particulate measures.

```
fit2 <- glm(DieselPM ~ TrafficPctl + Ozone, data = calenviroscreen)
summary(fit2)
Call:
glm(formula = DieselPM ~ TrafficPctl + Ozone, data = calenviroscreen)
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.23025068 0.01347754 17.08 <0.0000000000000002 ***
TrafficPctl 0.00355094 0.00009067 39.16 <0.0000000000000002 ***
Ozone -3.77418894 0.24967138 -15.12 <0.0000000000000002 ***
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for gaussian family taken to be 0.0545963)
    Null deviance: 537.24 on 7999 degrees of freedom
Residual deviance: 436.61 on 7997 degrees of freedom
   (35 observations deleted due to missingness)
AIC: -554.3
```
Number of Fisher Scoring iterations: 2

Linear regression: multiple predictors

Can also use tidy and glimpse to see the output nicely.

fit2 %>% tidy() %>% glimpse()

Rows: 3 Columns: 5 \$ term <chr> "(Intercept)" , "TrafficPctl" , "Ozone" \$ estimate <dbl> 0.23025068, 0.00355094, -3.77418894 \$ std.error <dbl> 0.01347753532, 0.00009067243, 0.24967137612 \$ statistic <dbl> 17.08403, 39.16229, -15.11663 \$ p.value <dbl> 0.00

Factors get special treatment in regression models - lowest level of the factor is the comparison group, and all other factors are **relative** to its values.

Let's create a variable that tells us whether a census tract has a high, middle, or low percentage of the population below the poverty line.

```
calenviroscreen <- calenviroscreen %>% mutate(
  PovertyPctl_level = case_when(
      Povertypctl > 0.75 \sim "high",
      <code>PovertyPctl</code> > 0.25 & <code>PovertyPctl</code> \leq 0.75 \sim "middle",
      PovertyPctl \leq 0.25 ~ "low",
     TRUE \sim NA
   \left( \frac{1}{2} \right)\left( \frac{1}{2} \right)
```
The comparison group that is not listed is treated as intercept. All other estimates are relative to the intercept.

```
fit3 <- glm(DieselPM ~ TrafficPctl + Ozone + factor(PovertyPctl_level), data = calenviroscreen)
summary(fit3)
```

```
Call:
qlm(formula = DieseIPM - TrafficPct1 + 0zone + factor(PovertPct1-level), data = calenviroscreen)
```
Coefficients:

```
 Estimate Std. Error t value
(Intercept) 0.22893847 0.01343002 17.047
TrafficPctl 0.00353551 0.00009034 39.137
0zone -3.73294307 0.24881820 -15.003
factor(PovertyPctl_level)low -0.09685048 0.05463573 -1.773
factor(PovertyPctl_level)middle -0.11329720 0.03672457 -3.085
                                    Pr(>\vert t \vert)(Intercept) < 0.0000000000000002 ***
TrafficPctl < 0.0000000000000002 ***
Ozone < 0.0000000000000002 ***
factor(PovertyPctl_level)low 0.07632 . 
factor(PovertyPctl_level)middle 0.00204 ** 
- - -Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for gaussian family taken to be 0.05357268)
    Null deviance: 523.61 on 7933 degrees of freedom
Residual deviance: 424.78 on 7929 degrees of freedom
  (101 observations deleted due to missingness)
AIC: -697.85
```
Number of Fisher Scoring iterations: 2

Relative to the level is not listed.

```
calenviroscreen <-
   calenviroscreen %>%
  mutate(PovertyPctl_level = factor(
    PovertyPctl_level,
   levels = c("low", "middle", "high") ))
fit4 <- glm(DieselPM ~ TrafficPctl + Ozone + PovertyPctl_level, data = calenviroscreen)
summary(fit4)
Call:
glm(formula = DieselPM ~ TrafficPctl + Ozone + PovertyPctl_level, 
    data = calenviroscreen)
Coefficients:
                        Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.13208799 0.05585244 2.365 0.0181 * 
TrafficPctl 0.00353551 0.00009034 39.137 <0.0000000000000002 ***
0zone -3.73294307  0.24881820 -15.003 <0.0000000000000002 ***
PovertyPctl_levelmiddle -0.01644672 0.06569819 -0.250 0.8023 
PovertyPctl_levelhigh 0.09685048 0.05463573 1.773 0.0763 . 
- - -Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for gaussian family taken to be 0.05357268)
    Null deviance: 523.61 on 7933 degrees of freedom
Residual deviance: 424.78 on 7929 degrees of freedom
   (101 observations deleted due to missingness)
AIC: -697.85
```
Number of Fisher Scoring iterations: 2

/

You can view estimates for the comparison group by removing the intercept in the GLM formula

 $y - x - 1$

Caveat is that the p-values change.

```
fit5 <- glm(DieselPM ~ TrafficPctl + Ozone + PovertyPctl_level - 1, data = calenviroscreen)
summary(fit5)
```

```
Call:
glm(formula = DieselPM ~ TrafficPctl + Ozone + PovertyPctl_level - 
     1, data = calenviroscreen)
```
Coefficients:

```
Estimate Std. Error t value Pr(>|t|)TrafficPctl 0.00353551 0.00009034 39.137 <0.0000000000000002 ***
0zone -3.73294307  0.24881820 -15.003 <0.0000000000000002 ***
PovertyPctl_levellow 0.13208799 0.05585244 2.365 0.0181 * 
PovertyPctl levelmiddle 0.11564127 0.03838198 3.013 0.0026 **
PovertyPctl_levelhigh 0.22893847 0.01343002 17.047 <0.0000000000000002 ***
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```
(Dispersion parameter for gaussian family taken to be 0.05357268)

 Null deviance: 919.65 on 7934 degrees of freedom Residual deviance: 424.78 on 7929 degrees of freedom (101 observations deleted due to missingness) AIC: -697.85

Number of Fisher Scoring iterations: 2

Linear regression: interactions

You can also specify interactions between variables in a formula $y - x1 + x2 + x1 * x2$. This allows for not only the intercepts between factors to differ, but also the slopes with regard to the interacting variable.

```
fit6 < -qlm(
 DieselPM ~ TrafficPctl + Ozone + PovertyPctl_level + TrafficPctl * PovertyPctl_level,
  data = calenviroscreen
)
tidy(fit6)
# A tibble: 7 \times 5 term estimate std.error statistic p.value
 <chr> <dbl> <dbl> <dbl> <dbl>
1 (Intercept) 0.200 0.113 1.77 7.62e- 2
2 TrafficPctl 0.00224 0.00186 1.21 2.28e- 1
3 Ozone -3.72 0.249 -14.9 7.90e-50
4 PovertyPctl_levelmiddle 0.00335 0.131 0.0256 9.80e- 1
5 PovertyPctl_levelhigh 0.0280 0.112 0.249 8.03e- 1
6 TrafficPctl:PovertyPctl_levelmiddle -0.000721 0.00227 -0.317 7.51e- 1
7 TrafficPctl:PovertyPctl_levelhigh 0.00131 0.00186 0.702 4.83e- 1
```
Linear regression: interactions

By default, ggplot with a factor added as a color will look include the interaction term. Notice the different intercept and slope of the lines.

```
ggplot(calenviroscreen, aes(x = DieselPM, y = TrafficPctl, color = PowertyPctl\_level)) +geom_point(size = 1, alpha = 0.1) +
  geom\_smooth(method = "glm", se = FALSE) + scale_color_manual(values = c("black", "grey45", "grey65", "grey85")) +
   theme_classic() +
  ylim(0, 100) +xlim(0, 3)
```


Generalized linear models (GLMs)

Generalized linear models (GLMs) allow for fitting regressions for noncontinuous/normal outcomes. Examples include: logistic regression, Poisson regression.

Add the family argument – a description of the error distribution and link function to be used in the model. These include:

- $binomial$ (link = "logit") outcome is binary **·**
- $poisson(link = "log")$ outcome is count or rate **·**
- others **·**

Very important to use the right test!

See this [case study](https://www.opencasestudies.org/ocs-bp-vaping-case-study/#Data_Analysis) for more information.

See ?family documentation for details of family functions.

Logistic regression

Let's look at a logistic regression example. We'll use the calenviroscreen dataset again. We will create a new binary variable based on the DieselPM percentile variable, so we can tell whether a census tract has high or low DieselPM emissions compared to the others.

```
calenviroscreen <-
   calenviroscreen %>%
     mutate(
       DieselPM_level = case_when
      (DieselPMPct1 > 0.75 ~- 1,DieselPMPctl \leq 0.75 \leq 0))
```
Logistic regression

Now that we've created the DieselPM level variable (where a 1 indicates the census tract is one of the top 75% when it comes to dieselPM emissions), we can run a logistic regression.

Let's explore how PovertyPctl level might predict DieselPM level.

```
# General format
qlm(y - x, data = DATASET NAME, family = binomial(link = "logit"))binom_fit <- glm(DieselPM_level ~ PovertyPctl_level, data = calenviroscreen, family = binomial(link = "logit"))
summary(binom_fit)
Call:
qlm(formula = DieselPM\_level - PowertyPct1\_level, family = binomial(link = "logit"), data = calenviroscreen)
Coefficients:
                               Estimate Std. Error z value Pr(>|z|)
(Intercept) 17.56606843873 932.48063065847 0.019 0.985
PovertyPctl_levelmiddle 0.00000004734 1118.59764091255 0.000 1.000
PovertyPctl_levelhigh -12.62378846430 932.48064030187 -0.014 0.989
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 666.77 on 7959 degrees of freedom
Residual deviance: 665.93 on 7957 degrees of freedom
   (75 observations deleted due to missingness)
AIC: 671.93
Number of Fisher Scoring iterations: 16
```
Logistic Regression

See this [case study](https://www.opencasestudies.org/ocs-bp-vaping-case-study/#Logistic_regression_%E2%80%9Cby_hand%E2%80%9D_and_by_model) for more information.

Odds ratios

An odds ratio (OR) is a measure of association between an exposure and an outcome. The OR represents the odds that an outcome will occur given a particular exposure, compared to the odds of the outcome occurring in the absence of that exposure.

Check out [this paper](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC2938757/).

Odds ratios

Use oddsratio(x, y) from the epitools() package to calculate odds ratios.

In this case, we're calculating the odds ratio for whether living in a high traffic area predicts high levels of DieselPM.

```
library(epitools)
calenviroscreen <-
   calenviroscreen %>%
     mutate(
       Traffic_level = case_when
      (TrafficPctl > 0.75 ~- 1,TrafficPct1 \leq 0.75 \leq 0)
```
response <- calenviroscreen %>% pull(DieselPM_level) predictor <- calenviroscreen %>% pull(Traffic_level)

Odds ratios

<code>Processing</code> math: 100% $|$ ased <code>estimate & mid-p</code> exact <code>CI"</code>

Use oddsratio(x, y) from the epitools() package to calculate odds ratios.

In this case, we're calculating the odds ratio for whether living in a high traffic area predicts high levels of DieselPM.

oddsratio(predictor, response) \$data Outcome Predictor 0 1 Total 0 23 37 60 1 35 7905 7940 Total 58 7942 8000 \$measure odds ratio with 95% C.I. Predictor estimate lower upper 0 1.0000 NA NA 1 139.3968 74.58837 260.5596 \$p.value two-sided Predictor midp.exact fisher.exact 0 NA NA 1 0 0.000000000000000000000000000000000007956334 two-sided Predictor 0 1 0.00 \$correction [1] FALSE attr(,"method")

/

Final note

Some final notes:

- Researcher's responsibility to **understand the statistical method** they use **·** underlying assumptions, correct interpretation of method results
- Researcher's responsibility to **understand the R software** they use meaning **·** of function's arguments and meaning of function's output elements

Summary

- glm() fits regression models: **·**
	- Use the $\tt{formula}$ = argument to specify the model (e.g., $y \sim x$ or $y \sim x1$
		- + x2 using column names)
	- Use data = to indicate the dataset **-**
	- Use family = to do a other regressions like logistic, Poisson and more **-**
	- summary() gives useful statistics **-**
- oddsratio() from the epitools package can calculate odds ratios (outside of logistic regression - which allows more than one explanatory variable) **·**
- **•** this is just the tip of the iceberg!

Resources (also on the [website!](https://daseh.org/resources.html))

For more check out:

- [this chapter](https://jhudatascience.org/tidyversecourse/model.html#linear-modeling) on modeling in this tidyverse book **·**
- [this chart on when to do what test](https://www.scribbr.com/statistics/statistical-tests/) **·**
- opencasestudies.org **·**

Content for similar topics as this course can also be found on Leanpub.

Lab Part 2

[Class Website](https://daseh.org/)

[Lab](https://daseh.org/modules/Statistics/lab/Statistics_Lab.Rmd)

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